Master Program in *Data Science and Business Informatics* Statistics for Data Science

Lesson 35 - Testing independence/association, Multiple sample testing of the mean

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Testing independence/association: discrete data

- Pearson's Chi-Square test of independence
- X and Y discrete (finite) distributions
- $(x_1, y_1) \dots, (x_n, y_n)$ bivariate observed dataset
- $H_0: X \perp Y \quad H_1: X \not\perp Y$
- Test statistic:

$$\chi^{2} = \sum_{i,j} \frac{(O_{i,j} - E_{i,j})^{2}}{E_{i,j}} = n \sum_{i,j} \frac{(O_{i,j}/n - p_{i,j}, p_{j,j})^{2}}{p_{i,j}, p_{j,j}} \sim \chi^{2}(df)$$

where $O_{i,j}$ is the number of observations of value X = i and Y = j, $E_{i,j} = np_{i,.}p_{.,j}$ where $p_{i,.} = \sum_j O_{i,j}/n$ and $p_{.,j} = \sum_i O_{i,j}/n$. $df = (n_x - 1)(n_y - 1)$ where n_x (resp., n_y) is the size of the support of X (resp., Y)

- Exact test when *n* is small: Fisher's exact test
- Paired data (e.g., before and after taking a drug): McNemar's test

Association between nominal variables: χ^2 -based

- Association measures based on Pearson χ^2
 - ▶ φ coefficient (or MCC, Matthews correlation coefficient)
 - $\hfill\square$ For 2 \times 2 contingency tables:

$$\phi = \sqrt{rac{\chi^2}{n}} \in [0,1]$$

► Cramer's V

 \Box For contingency tables larger than 2 \times 2:

$$V=\sqrt{rac{\chi^2}{n\cdot\min\left\{r-1,c-1
ight\}}}\in\left[0,1
ight]$$

where r and c are the number of rows and columns

Tschuprov's T

 $\hfill\square$ For contingency tables larger than 2 \times 2:

$$T = \sqrt{\frac{\chi^2}{n \cdot \sqrt{(r-1)(c-1)}}} \in [0,1]$$

where r and c are the number of rows and columns

See R script

[See [Lesson 16]

[Exercise. Show $\phi = |r_{xy}|$]

[sames as V if r = c]

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The G-test and Mutual Information

- G-test of independence
- X and Y discrete (finite) distributions
- $(x_1, y_1) \dots, (x_n, y_n)$ bivariate observed dataset
- $H_0: X \perp Y \quad H_1: X \not\perp Y$
- Test statistic:

$$G = 2\sum_{i,j} O_{i,j} \log \frac{O_{i,j}}{E_{i,j}} = 2\sum_{i,j} O_{i,j} \log \frac{O_{i,j}}{n \rho_{i,j} \rho_{i,j}} \sim \chi^2(df)$$

where $O_{i,j}$ is the number of observations of value X = i and Y = j, $E_{i,j} = np_{i,.}p_{.,j}$ where $p_{i,.} = \sum_j O_{i,j}/n$ and $p_{.,j} = \sum_i O_{i,j}/n$. $df = (n_x - 1)(n_y - 1)$ where n_x (resp., n_y) is the size of the support of X (resp., Y)

- Preferrable to Chi-Squared when numbers $(O_{ij} \text{ or } E_{ij})$ are small, asymptotically equivalent
- $G = 2 \cdot n \cdot I(O, E)$ where I(O, E) is the mutual information between O and E [See Lesson 16]

Testing correlation: continuous data

• Population correlation:

$$\rho = \frac{E[(X - \mu_X) \cdot (Y - \mu_Y)]}{\sigma_X \cdot \sigma_Y}$$

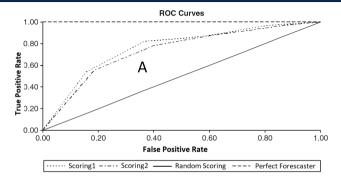
• Pearson's correlation coefficient:

$$r = \frac{\sum_{i=1}^{n} (x_i - \bar{x}) \cdot (y_i - \bar{y})}{\sqrt{\sum_{i=1}^{n} (x_i - \bar{x})^2 \cdot \sum_{i=1}^{n} (y_i - \bar{y})^2}}$$

- Assumption: joint distribution of X, Y is bivariate normal (or large sample)
- $(x_1, y_1) \dots, (x_n, y_n)$ bivariate observed dataset
- $H_0: \rho = 0$ $H_1: \rho \neq 0$
- Test statistics:

$$T=\frac{r\sqrt{n-2}}{\sqrt{1-r^2}}\sim t(n-2)$$

Testing AUC-ROC



- Binary classifier score $s_{ heta}(w) \in [0,1]$ where $s_{ heta}(w)$ estimate $\eta(w) = P_{\theta_{TRUE}}(C = 1|W = w)$
- ROC Curve
 - $TPR(p) = P(s_{\theta}(w) \ge p | C = 1)$ and $FPR(p) = P(s_{\theta}(w) | C = 0)$
 - ROC Curve is the scatter plot TPR(p) over FPR(p) for p ranging from 1 down to 0
 - AUC-ROC is the area below the curve What does AUC-ROC estimate?
 - ► Linearly related to Somer's D correlation index (a.k.a. Gini coefficient) [See Lesson 16]

Testing AUC-ROC

• AUC is the probability of correct identification of the order between two instances:

$$AUC = P_{\theta_{TRUE}}(s_{\theta}(W1) < s_{\theta}(W2) | C_{W1} = 0, C_{W2} = 1)$$

where (W1, C_{W1}) \sim f_{\theta_{TRUE}} and (W2, C_{W2}) \sim f_{\theta_{TRUE}}

- Related to $U = W \frac{n(n+1)}{2}$, where W is the Wilcoxon rank-sum test statistics [See Lesson 34]
- $s_{\theta}(W_1), \ldots, s_{\theta}(W_n) \sim F_{\theta_{TRUE}}|_{C=1}$ and $s_{\theta}(V_1), \ldots, s_{\theta}(V_m) \sim F_{\theta_{TRUE}}|_{C=0}$

$$U = \sum_{i=1}^{n} \sum_{j=1}^{m} S(s_{\theta}(W_i), s_{\theta}(V_j)) \qquad S(X, Y) = \begin{cases} 1 & \text{if } X > Y \\ \frac{1}{2} & \text{if } X = Y \\ 0 & \text{if } X < Y \end{cases}$$

- AUC-ROC = $U/(n \cdot m)$ is an estimator of AUC
- Normal approximation, DeLong's algorithm or bootstrap for confidence interval estimation

Omnibus tests and post-hoc tests

- $H_0: \theta_1 = \theta_2 = \ldots = \theta_k \ [= 0]$
- $H_1: \theta_i \neq \theta_j$ for some $i \neq j$
- Omnibus tests detect any of several possible differences
 - Advantage: no need to pre-specify which treatments are to be comparedand then no need to adjust for making multiple comparisons
- If H_1 is rejected (test significant), a *post-hoc test* to find which $\theta_i \neq \theta_j$
 - Everything to everything post-hoc compare all pairs
 - One to everything post-hoc compare a new population to all the others
- We distinguish a few cases:
 - Multiple linear regression (normal errors + homogeneity of variances, i.e., U_i ~ N(0, σ²)):
 Γ F-test + t-test
 - Equality of means (normal distributions + homogeneity of variances):
 - $\ \ \square \ \ ANOVA + Tukey/Dunnett$
 - Equality of means (general distributions):
 - \Box Friedman + Nemenyi

F-test for multiple linear regression

- $\boldsymbol{Y} = \boldsymbol{X} \cdot \boldsymbol{\beta} + \boldsymbol{U}$, where $\boldsymbol{Y} = (Y_1, \dots, Y_n)$, $\boldsymbol{U} = (U_1, \dots, U_n)$, and $\boldsymbol{X} = (\boldsymbol{x}_1, \dots, \boldsymbol{x}_n)$
 - $\boldsymbol{\beta}^{T} = (\alpha, \beta_1, \dots, \beta_k)$ and $\boldsymbol{x}_i = (1, x_i^1, \dots, x_i^k)$
 - Unexplained (residual) error $SSE = S(\beta) = \sum_{i=1}^{n} (y_i \mathbf{x}_i \cdot \beta)^2$
- Null model (or intercept-only model): $\mathbf{Y} = \mathbf{1} \cdot \alpha + \mathbf{U}$
 - ► Total error $SST = S(\alpha) = \sum_{i=1}^{n} (y_i \bar{y}_n)^2$ [residuals of the null model]
- Explained error $SSR = SST SSE = \sum_{i=1}^{n} (\bar{y}_n \boldsymbol{x}_i \cdot \boldsymbol{\beta})^2$
- Coefficient of determination $R^2 = SSR/SST = 1 SSE/SST$ [See Lesson 20]
 - Is the model useful? Fraction of explained error
- Is the model statistically significant?
- $H_0: \beta_1 = \ldots = \beta_k = 0$ $H_1: \beta_i \neq 0$ for all $i = 1, \ldots, k$
- Test statistic: $F = \frac{SSR}{SSE} \frac{n-k-1}{k} \sim F(k, n-k-1)$

Equality of means: ANOVA

- $H_0: \mu_1 = \mu_2 = \ldots = \mu_k$
- $H_1: \mu_1 \neq \mu_2$ for some $i \neq j$
- datasets $y_1^j, \ldots, y_{n_j}^j$ for $j = 1, \ldots, k$
 - ▶ responses of k 1 treatments and 1 control group
 - accuracies of k classifiers over $n_j = n$ datasets
- Linear regression model over dummy encoded *j*:

$$Y = \alpha + \beta_1 x_1 + \ldots + \beta_{k-1} x_{k-1}$$

- α = μ_k is the mean of the reference group (j = k)
 β_i = μ_i − μ_k
- F-test: $H_0: \beta_1 = \ldots = \beta_k = 0$, i.e., $\mu_j = \mu_k$
- Tukey HSD (Honest Significant Differences) is an all-pairs post-hoc test
- Dunnet test is a one-to-everything test

See R script

[one way ANOVA] [repeated measures/two way ANOVA]

Non-parametric test of equality of means: Friedman

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$$H_0: \mu_1 = \mu_2 = \ldots = \mu_k$$

- $H_1: \mu_1 \neq \mu_2$ for some $i \neq j$
- datasets x_1^j, \ldots, x_n^j for $j = 1, \ldots, k$
 - ► accuracies of k classifiers over n datasets
- Let r_i^j be the rank of x_i^j in x_i¹,...,x_i^k
 e.g., jth classifier w.r.t. ith dataset
- Average rank of classifier: $R_i = \frac{1}{n} \sum_{i=1}^{n} r_i^j$
- Under H_0 , we have $R_1 = \ldots = R_k$ and, for *n* and *k* large:

$$\chi_F^2 = \frac{12n}{k(k+1)} \left(\sum_{j=1}^k R_j^2 - \frac{k(k+1)^2}{4} \right) \sim \chi^2(k)$$

- Nemenyi test is an all-pairs post-hoc test
- Bonferroni correction is a one-to-everything test
- For unpaired observations, use Kruskal-Wallis test instead of Friedman test

See R script

[paired observations/repeated measures]

• On confidence intervals and statistical tests (with R code)

Myles Hollander, Douglas A. Wolfe, and Eric Chicken (2014) Nonparametric Statistical Methods. 3rd edition, John Wiley & Sons, Inc.